



Non recursive model for income data

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1. Introduction

In a previous example we showed how individual equations for a non-recursive system can be estimated separately using Two-Stage Least Squares (TSLS) and SIMPLIS syntax if a sufficient number of instrumental variables is available in the data. Here we illustrate how a non-recursive system can be estimated using LISREL syntax and maximum likelihood estimation. We again use the data from Gujarati (1995) to illustrate. Of interest here is the relationship between income and government spending.

The data given in **incomemoney.isf** are selected macro-economic data for the USA for the period 1970 to 1991. This file can be found in the **MVABOOK Examples\Chapter 2** folder.

The variables are:

- y_1 = income
- y_2 = money supply
- x_1 = investment expenditure
- x_2 = government spending on goods and services
- x_3 = interest rate on 6-month Treasury bills in %.

All variables except for x_3 are measured in billions of dollars.

	Y1	Y2	X1	X2	X3
1	10.11	6.28	1.50	2.08	6.56
2	10.97	7.17	1.75	2.24	4.51
3	12.07	8.05	2.06	2.49	4.47
4	13.50	8.61	2.43	2.70	7.18
5	14.59	9.09	2.46	3.06	7.93
6	15.86	10.23	2.26	3.64	6.12
7	17.68	11.64	2.86	3.93	5.27
8	19.74	12.87	3.58	4.26	5.51
9	22.33	13.89	4.34	4.69	7.57
10	24.89	14.97	4.80	5.20	10.02
11	27.08	16.29	4.68	6.13	11.37
12	30.31	17.93	5.58	6.98	13.78
13	31.50	19.52	5.03	7.71	11.08
14	34.05	21.86	5.47	8.40	8.75
15	37.77	23.74	7.19	8.93	9.80
16	40.39	25.69	7.14	9.70	7.66
17	42.69	28.11	7.18	10.28	6.03
18	45.40	29.11	7.49	10.66	6.05
19	49.00	30.71	7.94	11.09	6.92
20	52.51	32.27	8.32	11.82	8.04
21	55.22	33.39	7.99	12.74	7.47
22	56.77	34.40	7.21	13.33	5.49

2. Using LISREL syntax

The model previously considered was

$$\begin{aligned} y_1 &= \alpha_1 + \beta_1 y_2 + \gamma_1 x_1 + \gamma_2 x_2 + u_1 \\ y_2 &= \alpha_2 + \beta_2 y_1 + u_2 \end{aligned} \quad (1)$$

The model parameter matrices \mathbf{B} and $\mathbf{\Gamma}$ are

$$\mathbf{B} = \begin{bmatrix} 0 & \beta_1 \\ \beta_2 & 0 \end{bmatrix}, \quad \mathbf{\Gamma} = \begin{bmatrix} \gamma_1 & \gamma_2 \\ 0 & 0 \end{bmatrix}$$

Syntax is given in **incomemoney1b.lis**.

```

L incomemoney1b.lis
!Income and Money Supply Model Estimated by ML
!Using LISREL Syntax
da ni=5
ra=incomemoney.lsf
mo ny=2 nx=2 be=fi al=fr ka=fr
fr be(1,2) be(2,1)
fi ga(2,1) ga(2,2)
ou

```

Output for this analysis is as follows:

LISREL Estimates (Maximum Likelihood)

BETA

	Y1	Y2
Y1	- -	-2.165 (1.847) -1.173
Y2	0.614 (0.009) 67.794	- -

GAMMA

	X1	X2
Y1	1.504 (1.089) 1.381	8.703 (4.268) 2.039
Y2	- -	- -

Covariance Matrix of Y and X

	Y1	Y2	X1	X2
Y1	235.456			
Y2	144.314	88.836		
X1	34.575	21.244	5.407	
X2	56.749	34.868	8.324	13.758

Mean Vector of Eta-Variables

	Y1	Y2
	30.200	18.901

PHI

	X1	X2
X1	5.407 (1.710) 3.162	
X2	8.324 (2.680) 3.106	13.758 (4.351) 3.162

PSI

Note: This matrix is diagonal.

Y1	Y2
4.815	0.385
(4.930)	(0.122)
0.977	3.158

Squared Multiple Correlations for Structural Equations

Y1	Y2
0.996	0.999

Comparison of these results with the two other analyses of the same data shows that these results are different from those obtained using TSLS. TSLS works better than maximum likelihood in small samples, but ML has better large sample properties. While ML makes assumptions about multivariate normality, this may not hold for these data and TSLS, with less strong distributional assumptions, may be more appropriate.

Financial data also frequently tend to exhibit strong autocorrelation, in which case the assumption of independent observations become questionable. Multicollinearity also occurs frequently in data of this nature. Inspection of the covariance matrix for these data we find this to be the case here:

Income and Money Supply Model Estimated by ML

Covariance Matrix

	Y1	Y2	X1	X2
Y1	235.456			
Y2	144.314	88.836		
X1	34.531	21.264	5.407	
X2	56.757	34.865	8.324	13.758

Total Variance = 343.457 Generalized Variance = 1.706

Largest Eigenvalue = 342.789 Smallest Eigenvalue = 0.055

Condition Number = 78.631

WARNING: The Condition Number indicates severe multicollinearity.

One or more variables may be redundant.

The condition number is the square root of the ratio of the largest and smallest eigenvalues. A large condition number, as is observed here, indicates the presence of multicollinearity in the data. The warning message is printed by LISREL if this ratio exceeds 30.

3. Direct, indirect and total effects

In previous analyses, we considered the model defined in (1). Also assume that we have only y_2 and y_1 having an effect on each other as shown below:

$$y_1 \begin{matrix} \xrightarrow{\beta_{21}} \\ \xleftarrow{\beta_{12}} \end{matrix} y_2$$

Note that x_1 has no effect on y_2 . It does have an effect on y_1 though, and y_1 has an effect on y_2 . The same argument can be made about x_2 , which has no direct effect on y_1 but does have an effect on y_2 . As a result, it has an effect on y_1 .

Define a cycle for y_1 as one path to y_2 and one return to y_1 . The effect of one cycle on y_1 is thus $\beta_{21}\beta_{12}$. It is clear that multiple cycles would lead to the total effect on y_1 being the sum of the infinite geometric series

$$\beta_{21}\beta_{12} + (\beta_{21}\beta_{12})^2 + (\beta_{21}\beta_{12})^3 + \dots$$

which is $\beta_{21}\beta_{12} / (1 - \beta_{21}\beta_{12})$ if $\beta_{21}\beta_{12} < 1$.

Similarly, the general effect of \mathbf{y} on itself can then be written as

$$\mathbf{B} + \mathbf{B}^2 + \mathbf{B}^3 + \dots = (\mathbf{I} - \mathbf{B})^{-1} - \mathbf{I}$$

provided that the infinite series converges. The total effect of \mathbf{x} on \mathbf{y} can be written as

$$(\mathbf{I} + \mathbf{B} + \mathbf{B}^2 + \mathbf{B}^3 + \dots)\mathbf{\Gamma} = (\mathbf{I} - \mathbf{B})^{-1}\mathbf{\Gamma}.$$

In order to ensure stability of this system, convergence of the series is needed. A necessary and sufficient condition is that all the eigenvalues of \mathbf{B} are within the unit circle.

The table below summarizes the formulae for direct, indirect and total effects.

	$\mathbf{x} \rightarrow \mathbf{y}$	$\mathbf{y} \rightarrow \mathbf{y}$
Direct	$\mathbf{\Gamma}$	\mathbf{B}
Indirect	$(\mathbf{I} - \mathbf{B})^{-1}\mathbf{\Gamma} - \mathbf{\Gamma}$	$(\mathbf{I} - \mathbf{B})^{-1} - \mathbf{I} - \mathbf{B}$
Total	$(\mathbf{I} - \mathbf{B})^{-1}\mathbf{\Gamma}$	$(\mathbf{I} - \mathbf{B})^{-1} - \mathbf{I}$

To obtain estimates of the indirect and total effects and their standard errors in LISREL, the EF option should be added to the OU line. The following additional output is then available in the output file:

Total and Indirect Effects

Total Effects of X on Y		
	X1	X2
	-----	-----
Y1	0.645	3.734
	(0.330)	(0.209)

	1.956	17.873
Y2	0.396	2.295
	(0.203)	(0.128)
	1.956	17.953

Indirect Effects of X on Y

	X1	X2
	-----	-----
Y1	-0.858	-4.969
	(0.825)	(4.052)
	-1.041	-1.226
Y2	0.396	2.295
	(0.203)	(0.128)
	1.956	17.953

Total Effects of Y on Y

	Y1	Y2
	-----	-----
Y1	-0.571	-0.929
	(0.199)	(0.324)
	-2.863	-2.869
Y2	0.264	-0.571
	(0.122)	(0.199)
	2.156	-2.863

Largest Eigenvalue of $B \cdot B'$ (Stability Index) is 4.689

Indirect Effects of Y on Y

	Y1	Y2
	-----	-----
Y1	-0.571	1.236
	(0.199)	(1.437)
	-2.863	0.860
Y2	-0.351	-0.571
	(0.123)	(0.199)
	-2.851	-2.863

Under ML estimation, it seems as if the system is diverging as the Stability Index is larger than 1. However, if the stability index is calculated from the TSLS estimates instead, it is found to be 0.377.