

Estimating residuals using the Kenny-Judd data

1. Introduction

Here we use the Kenny-Judd data and estimate residuals for these data. The Kenny–Judd model is

$$y = \alpha + \gamma_1 \xi_1 + \gamma_2 \xi_2 + \gamma_3 \xi_1 \xi_2 + \zeta, \quad (1)$$

where ξ_1 and ξ_2 are latent variables and ζ is a random error term assumed to be uncorrelated with ξ_1 and ξ_2 . Kenny & Judd (1984) considered the case when there are two observable indicators x_1 and x_2 of ξ_1 and two observable indicators x_3 and x_4 of ξ_2 , such that

$$\begin{bmatrix} x_1 \\ x_2 \\ x_3 \\ x_4 \end{bmatrix} = \begin{bmatrix} \tau_1 \\ \tau_2 \\ \tau_3 \\ \tau_4 \end{bmatrix} + \begin{bmatrix} 1 & 0 \\ \lambda_2 & 0 \\ 0 & 1 \\ 0 & \lambda_4 \end{bmatrix} \begin{bmatrix} \xi_1 \\ \xi_2 \end{bmatrix} + \begin{bmatrix} \delta_1 \\ \delta_2 \\ \delta_3 \\ \delta_4 \end{bmatrix}. \quad (2)$$

Kenny & Judd (1984) did not include the constant intercept terms α and τ_i in (1) and (2), but as argued by Jöreskog & Yang (1996), these are necessary for a correct analysis of the model.

Solving for ξ_1 and ξ_2 in the first and third equation in (2) and substituting into (1) gives

$$y_1 = \alpha + \gamma_1 x_1 + \gamma_2 x_3 + \gamma_3 x_1 x_3 + u, \quad (3)$$

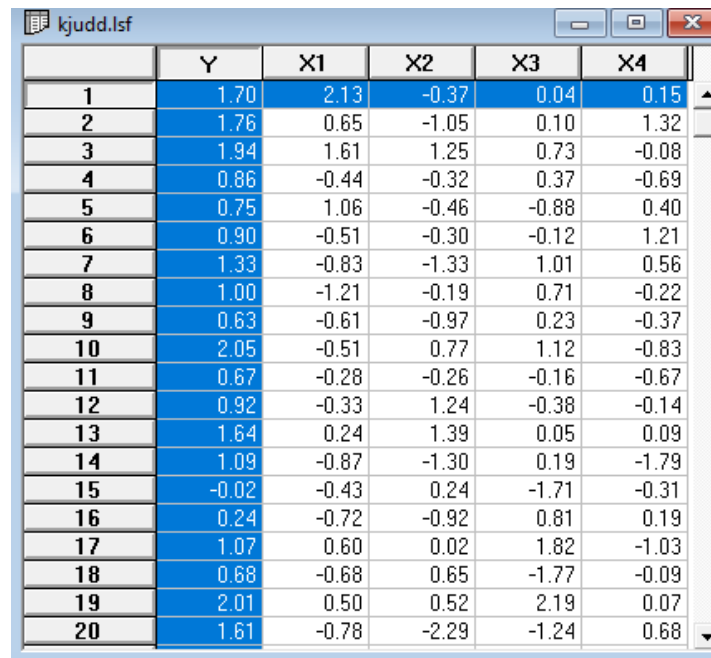
where

$$u = -\gamma_1 \delta_1 - \gamma_2 \delta_3 - \gamma_3 (x_1 \delta_3 + x_3 \delta_1 - \delta_1 \delta_3) + \zeta. \quad (4)$$

Note that the γ 's in (3) are the same as in (1) but the error term u is not the same as ζ . The error term u in (3) is correlated with the variables on the right side in (3) so that ordinary least squares cannot be used to estimate this equation. However,

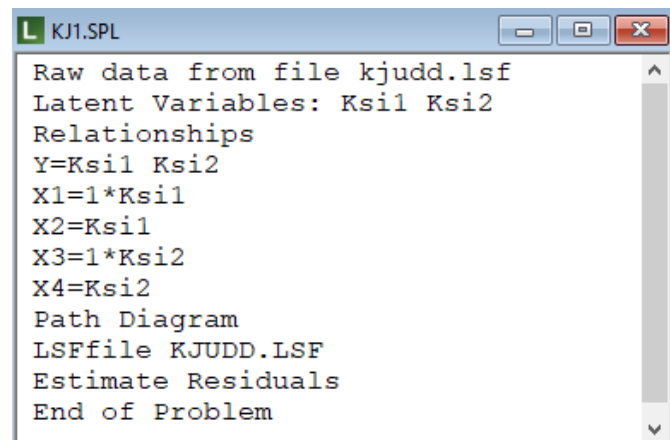
as shown by Bollen & Paxton (1998), x_2 , x_4 , and x_2x_4 can be used as instrumental variables. If the model holds, these instrumental variables are uncorrelated with u .

As a first step, we estimate the residuals using the data in the LSF file **kjudd.lsf**. The first few lines of this file are shown below.



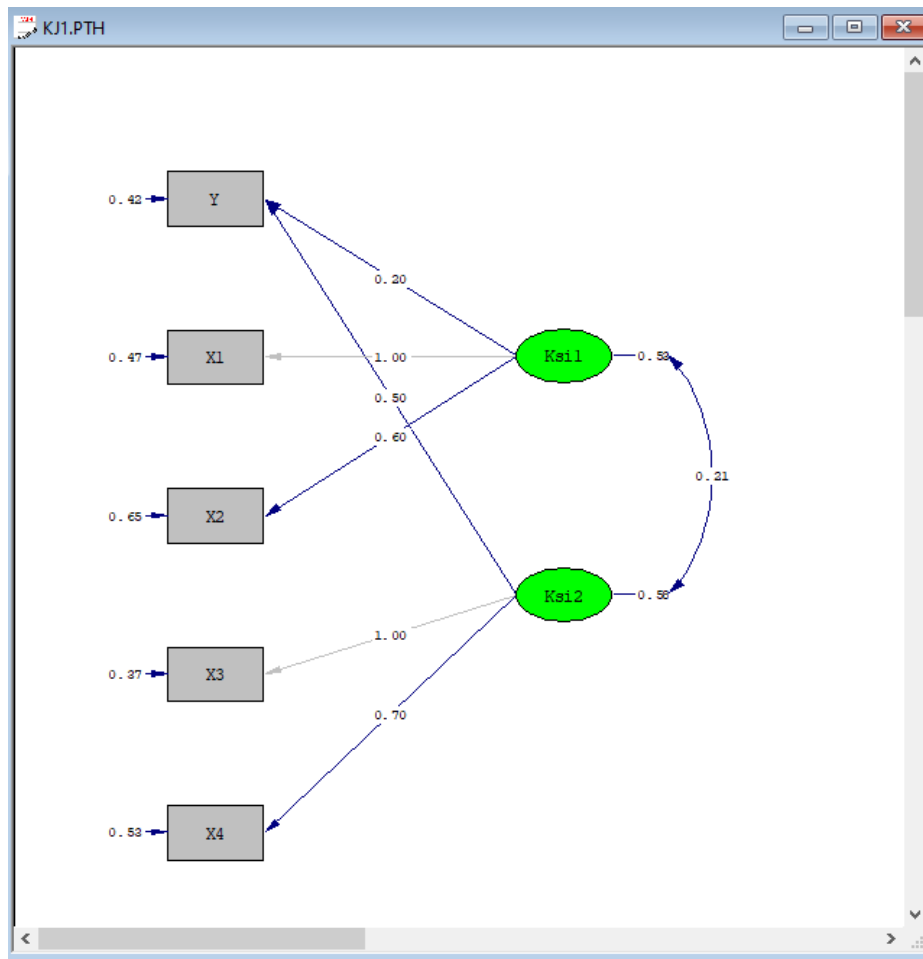
	Y	X1	X2	X3	X4
1	1.70	2.13	-0.37	0.04	0.15
2	1.76	0.65	-1.05	0.10	1.32
3	1.94	1.61	1.25	0.73	-0.08
4	0.86	-0.44	-0.32	0.37	-0.69
5	0.75	1.06	-0.46	-0.88	0.40
6	0.90	-0.51	-0.30	-0.12	1.21
7	1.33	-0.83	-1.33	1.01	0.56
8	1.00	-1.21	-0.19	0.71	-0.22
9	0.63	-0.61	-0.97	0.23	-0.37
10	2.05	-0.51	0.77	1.12	-0.83
11	0.67	-0.28	-0.26	-0.16	-0.67
12	0.92	-0.33	1.24	-0.38	-0.14
13	1.64	0.24	1.39	0.05	0.09
14	1.09	-0.87	-1.30	0.19	-1.79
15	-0.02	-0.43	0.24	-1.71	-0.31
16	0.24	-0.72	-0.92	0.81	0.19
17	1.07	0.60	0.02	1.82	-1.03
18	0.68	-0.68	0.65	-1.77	-0.09
19	2.01	0.50	0.52	2.19	0.07
20	1.61	-0.78	-2.29	-1.24	0.68

We use the syntax file **KJ1.SPL** to do so.



```
Raw data from file kjudd.lsf
Latent Variables: Ksi1 Ksi2
Relationships
Y=Ksi1 Ksi2
X1=1*Ksi1
X2=Ksi1
X3=1*Ksi2
X4=Ksi2
Path Diagram
LSFfile KJUDD.LSF
Estimate Residuals
End of Problem
```

The path diagram shows the estimates obtained under this model.



The maximum likelihood estimates for this model are given below.

LISREL Estimates (Maximum Likelihood)

Measurement Equations

Y = 0.204*Ksi1 + 0.500*Ksi2, Errorvar.= 0.424 , R² = 0.331
 Standerr (0.0560) (0.0552) (0.0237)
 Z-values 3.645 9.053 17.853
 P-values 0.000 0.000 0.000

X1 = 1.000*Ksi1, Errorvar.= 0.474 , R² = 0.526
 Standerr (0.0863)
 Z-values 5.496
 P-values 0.000

X2 = 0.599*Ksi1, Errorvar.= 0.646 , R² = 0.226
 Standerr (0.0974) (0.0417)
 Z-values 6.147 15.499
 P-values 0.000 0.000

X3 = 1.000*Ksi2, Errorvar.= 0.372 , R² = 0.609
 Standerr (0.0494)
 Z-values 7.523
 P-values 0.000

$X4 = 0.703 * Ksi2$, Errorvar.= 0.532 , $R^2 = 0.350$
 Standerr (0.0621) (0.0331)
 Z-values 11.331 16.055
 P-values 0.000 0.000

Covariance Matrix of Independent Variables

	Ksi1	Ksi2
Ksi1	0.527 (0.092) 5.698	
Ksi2	0.211 (0.030) 7.067	0.578 (0.061) 9.514

The LSF file **KJUDDnew.lsf** shown below is created as part of the output from this run. The variables R_X1 to R_X4 are the estimates of the measurement errors δ_1 to δ_4 , while R_Y represents the estimate of the measurement error ε_1 .

	Y	X1	X2	X3	X4	Ksi1	Ksi2	R_Y	R_X1	R_X2	R_X3	R_X4
1	1.70	2.13	-0.37	0.04	0.15	1.17	0.19	0.19	0.92	-1.07	-0.17	0.02
2	1.76	0.65	-1.05	0.10	1.32	0.17	0.52	0.28	0.44	-1.15	-0.44	0.96
3	1.94	1.61	1.25	0.73	-0.08	1.32	0.55	0.22	0.25	0.46	0.16	-0.46
4	0.86	-0.44	-0.32	0.37	-0.69	-0.40	-0.07	-0.20	-0.08	-0.08	0.42	-0.63
5	0.75	1.06	-0.46	-0.88	0.40	0.40	-0.46	-0.28	0.62	-0.70	-0.44	0.73
6	0.90	-0.51	-0.30	-0.12	1.21	-0.42	0.17	-0.28	-0.13	-0.04	-0.31	1.09
7	1.33	-0.83	-1.33	1.01	0.56	-0.80	0.69	-0.03	-0.07	-0.85	0.30	0.08
8	1.00	-1.21	-0.19	0.71	-0.22	-0.78	0.26	-0.15	-0.47	0.28	0.44	-0.39
9	0.63	-0.61	-0.97	0.23	-0.37	-0.69	-0.12	-0.35	0.04	-0.55	0.33	-0.28
10	2.05	-0.51	0.77	1.12	-0.83	0.00	0.57	0.58	-0.56	0.77	0.53	-1.22
11	0.67	-0.28	-0.26	-0.16	-0.67	-0.32	-0.39	-0.25	-0.00	-0.06	0.21	-0.39
12	0.92	-0.33	1.24	-0.38	-0.14	0.06	-0.30	-0.12	-0.43	1.21	-0.09	0.08
13	1.64	0.24	1.39	0.05	0.09	0.53	0.16	0.27	-0.33	1.08	-0.12	-0.01
14	1.09	-0.87	-1.30	0.19	-1.79	-0.86	-0.40	0.29	-0.05	-0.78	0.58	-1.50
15	-0.02	-0.43	0.24	-1.71	-0.31	-0.39	-1.27	-0.49	-0.08	0.48	-0.46	0.59
16	0.24	-0.72	-0.92	0.81	0.19	-0.78	0.24	-0.90	0.02	-0.45	0.55	0.03
17	1.07	0.60	0.02	1.82	-1.03	0.32	0.66	-0.51	0.24	-0.17	1.14	-1.49
18	0.68	-0.68	0.65	-1.77	-0.09	-0.33	-1.08	0.11	-0.39	0.85	-0.70	0.68
19	2.01	0.50	0.52	2.19	0.07	0.53	1.36	0.04	-0.07	0.21	0.81	-0.88
20	1.61	-0.78	-2.29	-1.24	0.68	-0.99	-0.40	0.83	0.17	-1.69	-0.86	0.97

Looking at the univariate summary statistics for these variables, obtained by using the **Data Screening option**, we see that the range of the measurement error R_X3 is the largest (-2.697; 2.310).

Univariate Summary Statistics for Continuous Variables

Variable	Mean	St. Dev.	Skewness	Kurtosis	Minimum	Freq.	Maximum	Freq.
Y	1.181	0.796	1.208	3.612	-1.290	1	5.680	1
X1	0.042	1.001	0.010	-0.097	-3.170	1	3.440	1
X2	-0.003	0.914	-0.137	0.074	-3.960	1	2.440	2
X3	0.018	0.975	0.014	0.019	-3.420	1	3.080	1
X4	-0.006	0.904	0.054	-0.011	-2.760	1	3.080	1
Ksi1	-0.000	0.726	0.074	0.064	-2.414	1	2.716	1
Ksi2	-0.000	0.760	0.419	0.286	-1.926	1	2.634	1
R_Y	-0.000	0.599	0.772	1.799	-1.738	1	2.828	1

R_X1	0.000	0.406	-0.004	0.194	-1.321	1	1.512	1
R_X2	0.000	0.731	-0.058	-0.044	-2.697	1	2.310	1
R_X3	0.000	0.399	-0.234	0.122	-1.653	1	1.139	1
R_X4	0.000	0.650	-0.033	0.010	-2.073	1	2.015	1

Next, we estimate the regression of Y on Ksi1, Ksi2 and Ksi1*Ksi2 using the syntax file **KJ2.PRL** and the newly created LSF file **KJUDDnew.lsf**.

```

Estimating the regression of Y on Ksi1, Ksi2 and Ksi1*Ksi2
SY=KJUDDnew.LSF
NE Ksi1Ksi2 = Ksi1*Ksi2
CO ALL
RG Y on Ksi1 Ksi2 Ksi1Ksi2
OU XU

```

The output for this analysis is as follows:

Estimated Equations

Y	=	1.085	+	0.198*Ksi1	+	0.482*Ksi2	+	0.458*Ksi1Ksi2
Standerr		(0.0174)		(0.0246)		(0.0242)		(0.0263)
t-values		62.380		8.032		19.918		17.410
P-values		0.000		0.000		0.000		0.000

+ Error, R² = 0.572

Error Variance = 0.272

The following chi-squares test the hypothesis that all regression coefficients are zero except the intercept.

Variable	-2lnL	Chi-square	df	Covariates
Y	1531.087	849.125	3	Ksi1 Ksi2 Ksi1Ksi2

Analysis of Variance Table

Regression d.f.	Residual d.f.	F	Covariates
362.072	3	270.687	996
		444.083	Ksi1 Ksi2 Ksi1Ksi2

We are also interested in estimating the regression of R_Y, the estimate of the measurement error ε_1 , on Ksi1*Ks2. The file **KJ3.PRL** is used for this purpose.



For this model we find that the estimated equations are

Estimated Equations

R_Y = - 0.0949 + 0.451*Ksi1Ksi2 + Error, R² = 0.244
 Standerr (0.0173) (0.0251)
 t-values -5.481 17.931
 P-values 0.000 0.000

Error Variance = 0.271

The following chi-squares test the hypothesis that all regression coefficients are zero except the intercept.

Variable	-2lnL	Chi-square	df	Covariates
R_Y	1531.959	279.265	1	Ksi1Ksi2

Analysis of Variance Table

Regression d.f.	Residual d.f.	F	Covariates
87.280 1	270.924 998	321.513	Ksi1Ksi2

From the estimated equation, we conclude that both the intercept and estimated coefficient of the interaction term is not significant.